

Lecture 12: Sept 10, 2013

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12.1 Introduction to Probability

Definition 12.1 *Let (Ω, Σ, P) be a finite probability space, where Ω is a sample space which is the set of all possible outcomes, Σ is a set of events where each event is a set containing zero or more outcomes, and P is the assignment of probabilities to the events. Let X and Y be a Random Variables (RV) on Ω . Then,*

1. *Expectation of X is $E[X] = \sum_{w \in \Omega} X[w]P(w)$*
2. *Variance of X is $Var[X] = E[(X - E[X])^2]$*
3. *Covariance of X and Y is $Cov[X, Y] = E[(X - E[X])(Y - E[Y])]$*

Variance measures the spread of a distribution.

Proposition 12.2 *If X, Y be RVs*

1. $E[X] = \sum_{x \in X(\Omega)} xP(X = x)$
2. $Var[X] = E[X^2] - E[X]^2$
3. $E[X + Y] = E[X] + E[Y]$
4. $Var[X + Y] = Var[X] + Var[Y] + 2 * Cov[X, Y]$
5. $Cov[X, Y] = E[XY] - E[X]E[Y]$

Theorem 12.3 *Let X, Y be independent RVs*

1. $E[XY] = E[X]E[Y]$
2. $Cov[X, Y] = 0$
3. $Var[X + Y] = Var[X] + Var[Y]$

Covariance is a weak measure of independence. Also, $Cov[X, Y] = 0$ does not imply that X, Y are independent.

Definition 12.4 *Let $P \in [0, 1]$ and let X be an RV*

1. X is p -bernoulli RV (or p -bernoulli distributed) if $X : \Omega \implies \{0, 1\}$ and

$$P(X = x) = \begin{cases} p & \text{if } x = 1 \\ 1 - p & \text{if } x = 0 \end{cases}$$

2. X is $B(n, p)$ binomially distributed if $X : \Omega \implies \{0, 1, \dots, n\}$ and

$$P(X = x) = \binom{n}{x} p^x (1 - p)^{1-x}$$

Proposition 12.5 *Examples of bernoulli and binomial distributions.*

- p -bernoulli models the flip of a p -biased coin
- Let X_1, \dots, X_n be p -bernoulli RVs. Define $X = \sum_{i=1}^n X_i$ and suppose X_i s are independent, then, X is $B(n, p)$ distributed

Theorem 12.6 *Let X be a RV.*

- If X is p -bernoulli, then
 $E[X] = p$ and $\text{Var}[X] = p(1 - p)$
- If X is $B(n, p)$, then
 $E[X] = np$ and $\text{Var}[X] = np(1 - p)$

Proof:

1. Let X be p -bernoulli.

(a)

$$\begin{aligned} E[X] &= 1 * p + 0 * (1 - p) \\ &= p \end{aligned}$$

(b)

$$\begin{aligned} \text{Var}[X] &= E[X^2] - E[X]^2 \\ &= E[X] - E[X]^2 \quad // \text{ since } X = X^2 \text{ for } p\text{-bernoulli} \\ &= p - p^2 \\ &= p(1 - p) \end{aligned}$$

2. Let X be $B(n, p)$.

(a)

$$\begin{aligned} E[X] &= E[\sum_{i=1}^n X_i] \\ &= \sum_{i=1}^n E[X_i] \quad // \text{ from proposition 1.12 (4)} \\ &= np \end{aligned}$$

(b)

$$\begin{aligned} \text{Var}[X] &= \text{Var}[\sum_{i=1}^n X_i] \quad // \text{ from proposition 1.15 (2)} \\ &= \sum_{i=1}^n \text{Var}[X_i] \quad // \text{ since } X_i \text{ are independent} \\ &= np(1 - p) \end{aligned}$$

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12.2 Concentration Inequalities

Theorem 12.7 Markov's inequality

Let X be an RV and $X \geq 0$, then,

1. For any $\lambda > 0$, $P(X > \lambda) \leq \frac{E[X]}{\lambda}$
2. For a monotone increasing function $g: \mathbb{R} \implies \mathbb{R}_+$, we have $P(X \geq \lambda) \leq \frac{E[g(x)]}{g(\lambda)}$

Theorem 12.8 Chebychev's inequality

Let X be an RV. For any $\lambda > 0$, we have, $P(|X - E[X]| \geq \lambda) \leq \frac{\text{Var}[X]}{\lambda^2}$

Theorem 12.9 Hoeffding's inequality

Let X_1, \dots, X_n be independent RVs with $a_k \leq X_k \leq b_k$ for all $k = 1, 2, \dots, n$. Let $X = \sum_{k=1}^n X_k$ and let $c_k = b_k - a_k$, then for any $\lambda > 0$,

1. $P(X - E[X] \geq \lambda) \leq e^{-\frac{2\lambda^2}{\sum_{k=1}^n c_k^2}}$
2. $P(X - E[X] \leq -\lambda) \leq e^{-\frac{2\lambda^2}{\sum_{k=1}^n c_k^2}}$
3. $P(|X - E[X]| \geq \lambda) \leq 2e^{-\frac{2\lambda^2}{\sum_{k=1}^n c_k^2}}$

Corollary 12.10 Chernoff bound

Let X_1, \dots, X_n be independent RVs with $P(X_k = 1) = p_k$ and $P(X_k = 0) = 1 - p_k$, $p_k \in [0, 1]$, and let $X = \sum_{k=1}^n X_k$. Then for any $\lambda > 0$, we have, $P(|X - E[X]| \geq \lambda) \leq 2e^{-\frac{2\lambda^2}{n}}$

Corollary 12.11 Let X_1, \dots, X_n be independent $\{-1, +1\}$ valued RVs with $P(X_k = +1) = P(X_k = -1) = 0.5$. Then for any $\lambda > 0$, we have, $P(|X| \geq \lambda) \leq 2e^{-\frac{\lambda^2}{2n}}$

Proof: Idea of proof of Hoeffding's theorem:

$$\begin{aligned}
 P(X - E[X] \geq \lambda) &= P(e^{t(X-E[X])} \geq e^{t\lambda}) \\
 &\leq e^{-t\lambda} E[e^{t(X-E[X])}] \quad // \text{Markov's inequality} \\
 &\leq e^{-t\lambda} e^{E[t(X-E[X])]} \quad // t \text{ is chosen to obtain the tightest bound}
 \end{aligned}$$

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